

Robust multichannel sparse recovery

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Menu

- 1 Introduction
- 2 Nonparametric sparse recovery
- 3 Simulation studies
- 4 Source localization with Sensor Arrays

Single Measurement Vector Model (SMV)

$$\mathbf{y}_{M \times 1} = \mathbf{\Phi}_{M \times N} \mathbf{x}_{N \times 1} + \mathbf{e}_{N \times 1}$$

- $\mathbf{\Phi} = (\boldsymbol{\phi}_1 \ \cdots \ \boldsymbol{\phi}_N) = (\boldsymbol{\phi}_{(1)} \ \cdots \ \boldsymbol{\phi}_{(M)})^H$,
 $M \times N$ known *measurement (design, system) matrix*
- $\mathbf{x} = (x_1, \dots, x_N)^T$, unobserved *signal vector*
- $\mathbf{e} = (e_1, \dots, e_M)^T$, random noise vector.

Objective

recover **K -sparse** or **compressible** signal from \mathbf{y} knowing $\mathbf{\Phi}$, K

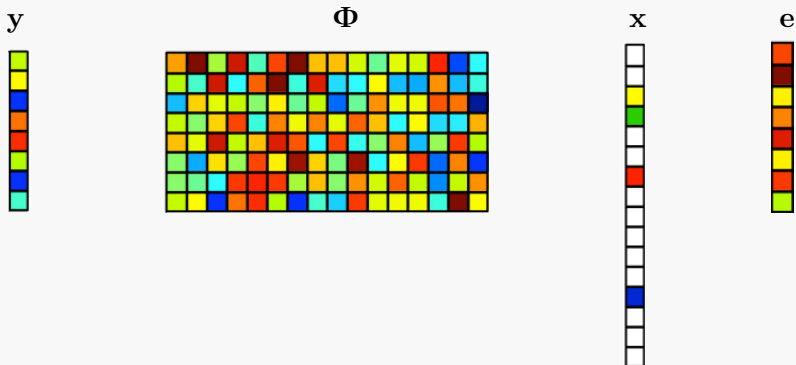
NOTE: typically, $M < N$ (ill-posed model) and $K \ll N$.

Single Measurement Vector Model (SMV)

$$\mathbf{y} = \Phi \mathbf{x} + \mathbf{e}$$

$M \times 1$ $M \times N$ $N \times 1$ $M \times 1$

- If we can determine $\Gamma = \text{supp}(\mathbf{x})$, then the problem reduces to conventional regression problem (more 'responses' than 'predictors'):



Multiple Measurement Vectors Model (MMV)

$$y_i = \Phi \mathbf{x}_i + \mathbf{e}_i, \quad i = 1, \dots, Q$$

In matrix form

$$\boxed{\begin{matrix} \mathbf{Y} & = & \Phi & \mathbf{X} & + & \mathbf{E} \\ M \times Q & & M \times N & N \times Q & & M \times 1 \end{matrix}}$$

- $\mathbf{Y} = (\mathbf{y}_1 \ \cdots \ \mathbf{y}_Q)$, observed **measurement matrix**
- $\mathbf{X} = (\mathbf{x}_1 \ \cdots \ \mathbf{x}_Q)$, unobserved **signal matrix**
- $\mathbf{E} = (\mathbf{e}_1 \ \cdots \ \mathbf{e}_Q)$, **noise matrix**

Objective

recover **K -sparse** (or **compressible**) signal matrix \mathbf{X} knowing \mathbf{Y} , Φ , K .

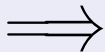
$N \times Q$ matrix \mathbf{X} is called **K -sparse** if $\|\mathbf{X}\|_0 = |\text{supp}(\mathbf{X})| \leq K$, where

$$\Gamma = \text{supp}(\mathbf{X}) = \bigcup_{i=1}^Q \text{supp}(\mathbf{x}_i) = \{j \in \{1, \dots, N\} : x_{jk} \neq 0 \text{ for some } k\}$$

Multiple Measurement Vectors Model (MMV)

Key idea:

signals are all predominantly supported on a common support set



Joint estimation can lead both to computational advantages and increased reconstruction accuracy

See [Tropp, 2006, Chen and Huo, 2006, Eldar and Rauhut, 2010, Duarte and Eldar, 2011, Blanchard et al., 2014].

Applications

- EEG/MEG [Gorodnitsky et al., 1995]
- equalization of sparse communications channels [Cotter and Rao, 2002]
- blind source separation [Gribonval and Zibulevsky, 2010]
- source localization using sensor arrays [Malioutov et al., 2005]
- ...

Matrix norms

The **mixed $\ell_{p,q}$ norm** of \mathbf{X} for $p, q \in [1, \infty)$

$$\|\mathbf{X}\|_{p,q} = \left(\sum_i \left(\sum_j |x_{ij}|^p \right)^{q/p} \right)^{1/q} = \left(\sum_i \|\mathbf{x}_{(i)}\|_p^q \right)^{1/q}.$$

- If $p = q \implies$ matrix p -norm $\|\mathbf{X}\|_{p,p} = \|\mathbf{X}\|_p$.
- ℓ_2 (Frobenius) norm $\|\cdot\|_2$ is denoted shortly as $\|\cdot\|$.

The **mixed ℓ_∞ -norms** of $\mathbf{X} \in \mathbb{C}^{N \times Q}$

- $\|\mathbf{X}\|_{p,\infty} = \max_{i \in \{1, \dots, N\}} \|\mathbf{x}_{(i)}\|_p$
- $\|\mathbf{X}\|_{\infty,q} = \left(\sum_i (\max_{j \in \{1, \dots, Q\}} |x_{ij}|)^q \right)^{1/q}$.

The row **ℓ_0 quasi-norm** of \mathbf{X}

$$\|\mathbf{X}\|_0 = |\text{supp}(\mathbf{X})| = \text{nr. of number of nonzero rows of } \mathbf{X}$$

Optimization problem

$$\min_{\mathbf{X}} \|\mathbf{Y} - \Phi\mathbf{X}\|^2 \quad \text{subject to} \quad \|\mathbf{X}\|_0 \leq K$$

is combinatorial (NP-hard). More feasible approaches:

- Optimization (Convex relaxation)
- Greedy pursuit:
 - Orthogonal Matching Pursuit (OMP)
 - Normalized Iterative Hard Thresholding (NIHT)
 - Normalized Hard Thresholding Pursuit (NHTP)
 - Compressive Sampling MP (CoSaMP)
- Guaranteed to perform very well under suitable conditions (restricted isometry property on Φ , non impulsive noise \mathbf{e})
- NIHT is simple and scalable, convenient for problems where the support is of primary interest

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Convex relaxation

IDEA: replace $\|\mathbf{X}\|_0$ by $\ell_{p,1}$ norm

- [Tropp, 2006] uses $p = \infty$:

$$\min_{\mathbf{X}} \|\mathbf{Y} - \Phi \mathbf{X}\|^2 + \lambda \|\mathbf{X}\|_{\infty,1}$$

where $\|\mathbf{X}\|_{\infty,1} = \sum_i \|\mathbf{x}_{(i)}\|_{\infty}$.

- [Malioutov et al., 2005] uses $p = 2$:

$$\min_{\mathbf{X}} \|\mathbf{Y} - \Phi \mathbf{X}\|^2 + \lambda \|\mathbf{X}\|_{2,1}$$

where $\|\mathbf{X}\|_{2,1} = \sum_i \|\mathbf{x}_{(i)}\|$.

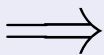
- ☹ good recovery depends on the choice of regularization parameter $\lambda > 0$
- ☹ not robust due to the used data fidelity term (LS-loss)
- ☹ not as scalable as greedy methods

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Nonparametric approach based on mixed norms

$$\min_{\mathbf{X}} \underbrace{\|\mathbf{Y} - \Phi\mathbf{X}\|_{p,q}^q}_{\text{data fidelity}} \quad \text{subject to} \quad \underbrace{\|\mathbf{X}\|_0}_{\text{sparsity constraint}} \leq K \quad (P_{p,q})$$



mixed ℓ_1 norms provide robustness, i.e., give larger weights on small residuals and less weight on large residuals.

We consider the cases

- $(p, q) = (1, 1)$: $\|\mathbf{Y} - \Phi\mathbf{X}\|_1 = \sum_i \sum_j |y_{ij} - \Phi_{(i)}^H \mathbf{x}_j|$
robust norm for errors i.i.d. in space and time
- $(p, q) = (2, 1)$: $\|\mathbf{Y} - \Phi\mathbf{X}\|_{2,1} = \sum_i \|\mathbf{y}_{(i)} - \mathbf{X}^H \phi_{(i)}\|$
robust norm for errors i.i.d. in "space" only

and devise greedy simultaneous NIHT (SNIHT) algorithm

Normalized iterative hard thresholding (NIHT)

Conventional $\ell_{2,2}$ approach :

$$\mathbf{X}^{n+1} = \underset{\substack{\uparrow \\ \text{hard thresholding}}}{H_K} \left(\mathbf{X}^n + \overset{\text{stepsize}}{\mu^{n+1}} \underbrace{\Phi^H(\mathbf{Y} - \Phi \mathbf{X}^n)}_{\substack{\uparrow \\ \text{neg. gradient } \nabla_{\mathbf{X}^*} \|\cdot\|^2}} \right)$$

- $\psi_{2,2}(\mathbf{X}) = \mathbf{X} \implies$ Conventional $\ell_{2,2}$ approach
- $\psi_{1,1}(\mathbf{X}) = [\text{sign}(x_{ij})]$ (elementwise operation of $S(\cdot)$)
- $\psi_{2,1}(\mathbf{X}) = [\text{sign}(x_{(j)})]$ (rowwise operation of $S(\cdot)$)

Normalized iterative hard thresholding (NIHT)

In our mixed norm $\ell_{p,q}$ approach:

$$\mathbf{X}^{n+1} = \underset{\substack{\uparrow \\ \text{hard thresholding}}}{H_K} \left(\mathbf{X}^n + \overset{\text{stepsize}}{\mu^{n+1}} \underbrace{\Phi^H \psi_{p,q}(\mathbf{Y} - \Phi \mathbf{X}^n)}_{\substack{\uparrow \\ \text{neg. gradient } \nabla_{\mathbf{X}^*} \|\cdot\|_{p,q}^q}} \right)$$

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Spatial sign

$$\text{sign}(\mathbf{x}) = \begin{cases} \mathbf{x}/\|\mathbf{x}\|, & \text{for } \mathbf{x} \neq 0 \\ 0, & \text{for } \mathbf{x} = 0 \end{cases}$$

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SNIHT(p, q) algorithm

input : \mathbf{Y} , Φ , sparsity K , mixed norm indices (p, q)

output : $(\mathbf{X}^{n+1}, \Gamma^{n+1})$ as estimates of \mathbf{X} and $\Gamma = \text{supp}(\mathbf{X})$

initialize: $\mathbf{X}^0 = \mathbf{0}$, $\mu^0 = 0$, $n = 0$, $\Gamma^0 = \emptyset$

1 $\Gamma^0 = \text{supp}(H_K(\Phi^H \psi_{p,q}(\mathbf{Y})))$

while *halting criterion false* **do**

2 $\mathbf{R}_\psi^n = \psi_{p,q}(\mathbf{Y} - \Phi \mathbf{X}^n)$

3 $\mathbf{G}^n = \Phi^H \mathbf{R}_\psi^n$

4 $\mu^{n+1} = \text{CompStepsize}(\Phi, \mathbf{G}^n, \Gamma^n, \mu^n, p, q)$

5 $\mathbf{X}^{n+1} = H_K(\mathbf{X}^n + \mu^{n+1} \mathbf{G}^n)$

6 $\Gamma^{n+1} = \text{supp}(\mathbf{X}^{n+1})$

7 $n = n + 1$

end

Computation of the step size μ^{n+1}

- For convergence, it is important that the stepsize is adaptively controlled at each iteration.
- Given the found support Γ^n is correct, one can choose μ^{n+1} as the min. of the objective fnc for the gradient ascent direction $\mathbf{X}^n + \mu \mathbf{G}^n|_{\Gamma^n}$:

$$L(\mu) = \|\mathbf{Y} - \Phi(\mathbf{X}^n + \mu \mathbf{G}^n|_{\Gamma^n})\|_{p,q}^q = \|\mathbf{R}^n - \mu \mathbf{B}\|_{p,q}^q$$

where $\mathbf{R}^n = \mathbf{Y} - \Phi \mathbf{X}^n$ and $\mathbf{B} = \Phi \mathbf{G}^n|_{\Gamma^n}$

- For $p = q$: simple linear regression problem based on ℓ_p -norm, with response $\mathbf{r} = \text{vec}(\mathbf{R}^n)$ and predictor $\mathbf{b} = \text{vec}(\mathbf{B})$.

$$\Rightarrow \mu^{n+1} = \|\mathbf{G}^n_{(\Gamma^n)}\|^2 / \|\Phi \mathbf{G}^n_{(\Gamma^n)}\|^2 \text{ for } p = q = 2$$

- For $p = q = 1$ and $(p, q) = (2, 1)$, minimizer of $L(\mu)$ can not be found in closed-form.

Computation of the step size μ^{n+1}

- When $p = q = 1$, the solution μ verifies a fixed point (FP) equation

$$\mu = H(\mu)$$

$$H(\mu) = \left(\sum_{ij} |\tilde{r}_{ij}|^{-1} |b_{ij}|^2 \right)^{-1} \sum_{ij} |\tilde{r}_{ij}|^{-1} \operatorname{Re}(b_{ij}^* r_{ij}),$$

and $\tilde{\mathbf{R}} = \mathbf{R}^n - \mu \mathbf{B}$ (depends on unknown μ) and $\mathbf{B} = \Phi \mathbf{G}_{(\Gamma^n)}$.

- We select μ^{n+1} as the **1-step FP iterate**

$$\mu^{n+1} = H(\mu^n),$$

where μ^n is the value of the stepsize at previous n th iteration.

- Often this approximation was within $\sim 1e^{-3}$ to minimizer of $L(\mu)$.
- Same approach is used in the case $(p, q) = (2, 1)$.

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Simulation study

$$\mathbf{Y} = \mathbf{\Phi} \mathbf{X} + \mathbf{E}$$

$M \times Q$ $M \times N$ $N \times Q$ $M \times 1$

Set-up

- Measurement matrix $\mathbf{\Phi}$: $\phi_{ij} \sim \mathcal{CN}(0, 1)$ with unit-norm columns
- $\Gamma = \text{supp}(\mathbf{X})$ randomly drawn K elements from $\{1, \dots, N\}$
- signal matrix \mathbf{X} : $|x_{ij}| = 10$ and $\text{Arg}(x_{ij}) \sim \text{Unif}(0, 2\pi) \forall i \in \Gamma$.

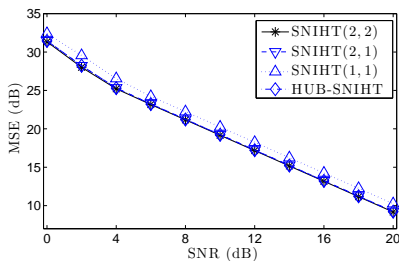
For $\ell = 1, \dots, L$ MC-trials ($L = 1000$), compute $\mathbf{X}^{[\ell]}$ and $\Gamma^{[\ell]}$

Performance measures

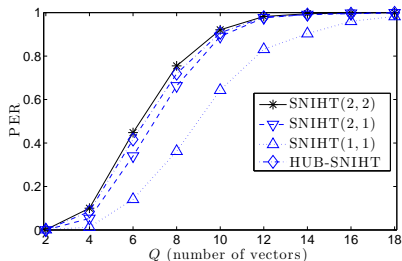
- mean squared error $\text{MSE}(\hat{\mathbf{X}}) = \frac{1}{LQ} \sum_{\ell=1}^L \|\hat{\mathbf{X}}^{[\ell]} - \mathbf{X}^{[\ell]}\|^2$.
- probability of exact recovery $\text{PER} = \frac{1}{L} \sum_{\ell=1}^L \mathbb{I}(\hat{\Gamma}^{[\ell]} = \Gamma^{[\ell]})$.

IID Gaussian noise, $e_{ij} \sim \mathcal{CN}(0, \sigma^2)$

MSE vs SNR for $Q=16$



PER vs Q at $\text{SNR} = 5$ dB

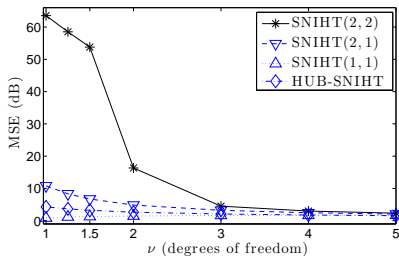


SNIHT(2,2)

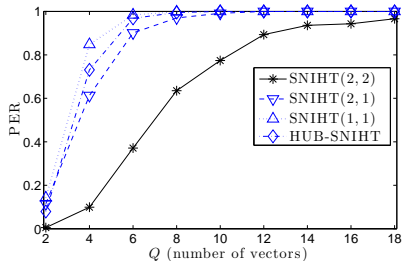
but SNIHT(2,1) and HUB-SNIHT (threshold $c = 1.517$) loose only 0.09 dB in MSE ($Q = 16$)

IID t -distributed noise, $e_{ij} \sim \mathcal{C}t_{\nu}(0, \sigma^2)$, $\sigma^2 = \text{Med}_F(|e_i|^2)$

SNR(σ) = 30 dB

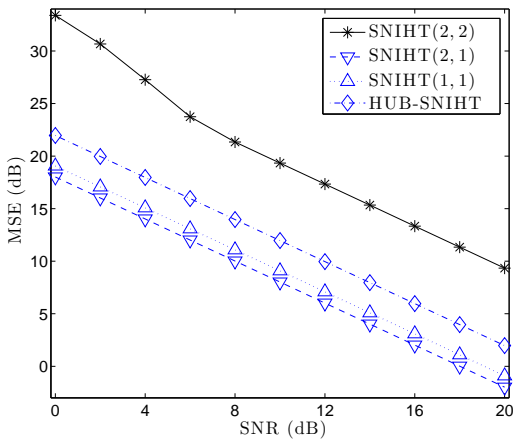


SNR(σ) = 10 dB, $\nu = 3$



SNIHT(1, 1)

Row IID compound Gaussian inverse Gaussian texture, $e_{(i)} \sim \text{CIG}_\lambda(0, \sigma^2 \mathbf{I}_Q)$



SNIHT(2,1)

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Sensor array and narrowband, farfield source model

- M sensors, K sources ($M > K$):
- Model for sensor measurements at time instant t

$$\mathbf{y}(t) = \mathbf{A}(\boldsymbol{\theta})\mathbf{x}(t) + \mathbf{e}(t),$$

- steering matrix $\mathbf{A}(\boldsymbol{\theta}) = (\mathbf{a}(\theta_1) \ \dots \ \mathbf{a}(\theta_K))$
- unknown (distinct) **Direction of Arrivals (DOA's)** $\theta_1, \dots, \theta_K$
- known array manifold $\mathbf{a}(\theta)$ (e.g., ULA)
- unknown signal waveforms $\mathbf{x}(t) = (x_1(t), \dots, x_K(t))^T$

Objective

Estimate the DOA's of the sources given the snapshots $\mathbf{y}(t_1), \dots, \mathbf{y}(t_Q)$ and the number of sources K impinging on the sensor array.

Multichannel sparse recovery approach to source localization

- Construct an $M \times N$ overcomplete steering matrix $\mathbf{A}(\tilde{\boldsymbol{\theta}})$ for a grid of all source locations $\tilde{\boldsymbol{\theta}}_1, \dots, \tilde{\boldsymbol{\theta}}_N$ of interest.
- Suppose that $\tilde{\boldsymbol{\theta}}$ contains the true DOA's to some accuracy.
- The array model in matrix form then rewrites as

$$\mathbf{Y} = (\mathbf{y}(t_1) \quad \cdots \quad \mathbf{y}(t_Q)) = \mathbf{A}(\tilde{\boldsymbol{\theta}})\mathbf{X} + \mathbf{E},$$

where $\mathbf{X} \in \mathbb{C}^{N \times Q}$ is K -sparse, with K non-zero rows containing the source waveforms at time instants t_1, \dots, t_Q .

$\implies K$ -sparse MMV model

- In the multichannel sparse recovery formulation $\mathbf{A}(\tilde{\boldsymbol{\theta}})$ is completely known, and we can use SNIHT methods to identify the support.

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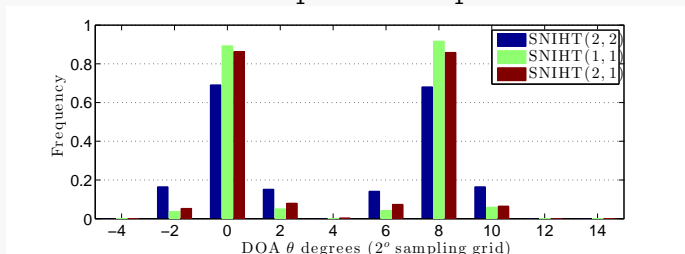
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Simul set-up

- ULA of $M = 20$ sensors with half a wavelength interelement spacing.
- $K = 2$ independent Gaussian sources from $\theta_1 = 0^\circ$ and $\theta_2 = 8^\circ$.
- Temporally/spatially white Gaussian error terms.
- Low SNR = 0 dB.
- Uniform grid $\tilde{\theta}$ on $[-90, 90]$ with 2° degree spacing
- We compute PER rates and relative frequencies of found DOA estimates in the grid for 1000 Monte Carlo runs.

Number of snapshots $Q = 4$

DOA's at $\theta_1 = 0^\circ$ and $\theta_1 = 8^\circ$.



PER rates

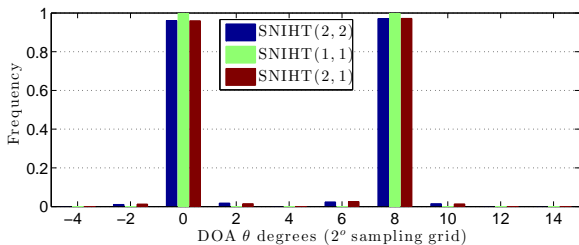
- 1 SNIHT(1, 1) = 81%
- 2 SNIHT(2, 1) = 73.1%
- 3 HUB-SNIHT = 42.6%
- 4 SNIHT = 37.6%



SNIHT(1, 1)

Number of snapshots $Q = 40$

DOA's at $\theta_1 = 0^\circ$ and $\theta_1 = 8^\circ$.



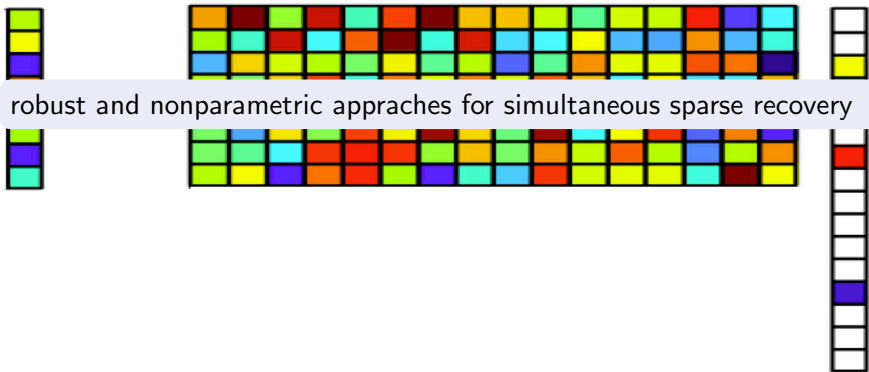
PER rates

- 1 SNIHT(1,1) = 100%
- 2 SNIHT(2,1) = 73.1%
- 3 HUB-SNIHT = 42.6%
- 4 SNIHT = 37.6%



SNIHT(1,1)

Conclusions



Conclusions

robust and nonparametric approaches for simultaneous sparse recovery



Different method, SNIHT(1, 1), SNIHT(2, 1), HUB-SNIHT, can be selected for a problem at hand (iid/correlated noise, high-breakdown, etc)



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Fast and scalable greedy SNIHT algorithm: **increase in robustness does not imply increase in computation time!**

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robust and nonparametric approaches for simultaneous sparse recovery



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Fast and scalable greedy SNIHT algorithm: **increase in robustness does not imply increase in computation time!**

Applications in source localization and image denoising, etc.

This talk based on



E. Ollila (2015a)

Nonparametric Simultaneous Sparse Recovery: an Application to Source Localization

PDF (ArXiv): <http://arxiv.org/pdf/1502.02441v1>
EUSIPCO'15, submitted.



E. Ollila (2015b).

Robust Simultaneous Sparse Approximation

A Festschrift in Honor of Hannu Oja, Springer, Sept. 2015.



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